### Image restoration

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Cours de Master M2

Première partie I

Image degradations

### Image degradations - Noise



**Noise Sources :** *Shot noise* (discrete nature of light), readout noise, transmission errors, etc.

## Image degradations - Blur



Motion blur. Source : flicker

## Image degradations - Blur



Optical blur. Wikimedia Commons

## Image degradations - Missing pixels



Source : JD.

### Original image



Additive Gaussian noise with  $\sigma = 20$  (image range in [0, 255])



#### Poisson noise



Impulse noise with p = 0.3



#### Gaussian blur



#### Motion blur



Random missing pixels



#### Missing region



Deuxième partie II

Global methods

$$\widetilde{U} = AU + N$$

- A linear degradation, known
- *U* original image  $\in \mathbb{R}^{n \times m}$
- N noise

 $\widetilde{U} = AU + N$ 

- A linear degradation, known : blur (deconvolution), missing pixels (inpainting or superresolution)
- U original image  $\in \mathbb{R}^{n \times m}$
- N noise

 $\widetilde{U} = AU + N$ 

- A linear degradation, known : blur (deconvolution), missing pixels (inpainting or superresolution)
- *U* original image ∈ ℝ<sup>n×m</sup> : an *a priori* on *U* is eventually known (regularity, content)
- N noise

 $\widetilde{U} = AU + N$ 

- A linear degradation, known : blur (deconvolution), missing pixels (inpainting or superresolution)
- *U* original image ∈ ℝ<sup>n×m</sup> : an *a priori* on *U* is eventually known (regularity, content)
- N noise :  $N = h(U)\varepsilon$ , the function h and the distribution of  $\varepsilon$  are known.

### Direct solution

$$\widehat{U} = \arg\min_{U} \|\widetilde{U} - AU\|^{2}$$
$$\Rightarrow$$
$$A^{*}\widetilde{U} - A^{*}A\widehat{U} = 0$$

where  $A^*$  is the adjoint of A.

The problem is **ill-posed** since  $A^*A$  is not always one-to-one and its eigenvalues may be small.

The operator  $D : \mathbb{R}^{n \times m} \to \mathbb{R}^{n \times m}$  which minimizes the risk  $\mathbb{E}[||U - D\widetilde{U}||^2].$ 

is the conditionnal expectation

 $D\widetilde{U} = \mathbb{E}[U/\widetilde{U}].$ 

D is **not necessarily linear**, and is quite difficult to estimate in practice, even if one knows the laws of the noise N and the image U (*a priori* law).

### Wiener

Assume that  $\mathbb{E}[N] = 0$  (the noise is centered) and  $\mathbb{E}[U] = 0$  (if not, replace U by  $U - \mathbb{E}[U]$ ).

The **linear** operator D minimizing  $\mathbb{E}[||U - D\widetilde{U}||^2]$  is the Wiener filter

 $D = \mathbb{E}[U\widetilde{U}^{\mathsf{T}}] \cdot \mathbb{E}[\widetilde{U}\widetilde{U}^{\mathsf{T}}]^{-1} = \Sigma_{U} \cdot A^{\mathsf{T}} (A\Sigma_{U}A^{\mathsf{T}} + \Sigma_{N})^{-1}.$ 

### Wiener deconvolution

Assume that

- A is a convolution with a discrete filter a
- the noise N is independent of the signal U.

We look for a convolution filter g which minimizes

$$\mathbb{E}[\|\boldsymbol{g}\star\widetilde{\boldsymbol{U}}-\boldsymbol{U}\|^2].$$

Then, in the Fourier domain

$$\widehat{g} = rac{\widehat{a}^*}{rac{\mathbb{E}[|\widehat{N}|^2]}{\mathbb{E}[|\widehat{U}|^2]} + |\widehat{a}|^2},$$

where  $\hat{a}$  is the discrete fourier transform (DFT) of a. In practice, we do not have access to  $\mathbb{E}[|\hat{U}|^2]$ ...

## Wiener



## Wiener



## Maximum *a posteriori*

Replace 
$$\mathbb{E}[U/\widetilde{U}]$$
 by  
 $\arg \max_{U} \mathbb{P}[U/\widetilde{U}] = \arg \max_{U} \mathbb{P}[\widetilde{U}/U] \mathbb{P}[U]$   
 $= \arg \min_{U} -\log \mathbb{P}[\widetilde{U}/U] - \log \mathbb{P}[U]$   
 $= \arg \min_{U} \underbrace{G(U,\widetilde{U})}_{\text{data fidelity term}} + \underbrace{F(U)}_{a \text{ priori on } U}$ .

### Maximum a posteriori

Replace 
$$\mathbb{E}[U/\widetilde{U}]$$
 by  
 $\arg \max_{U} \mathbb{P}[U/\widetilde{U}] = \arg \max_{U} \mathbb{P}[\widetilde{U}/U] \mathbb{P}[U]$   
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### Example

$$N \sim \mathcal{N}(0, \sigma^2 Id)$$
,  $p_U \propto e^{-F(u)}$ , then

$$rgmin_U \|\widetilde{U} - AU\|_2^2 + F(U),$$

with  $||||_2$  the euclidean norm on  $\mathbb{R}^{n \times m}$ .

### Tykhonov regularization

[Tikhonov, Arsenin, 1977]

$$E(u) = \frac{1}{2} \|\widetilde{u} - Au\|_2^2 + \lambda \int |\nabla u|^2,$$

where |
abla u| denotes the norm of abla u in  $\mathbb{R}^2$ 

### Tykhonov regularization

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Minimization : convex, differentiable, gradient descent with

$$\frac{\partial E}{\partial u} = A^* (Au - \tilde{u}) - 2\lambda \Delta u.$$

Or solve  $A^*(Au - \tilde{u}) - 2\lambda\Delta u = 0$  using the discrete Fourier transform.

# Tykhonov for denoising

$$A = Id$$
,  $N \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma = 0.2$  (image range = [0, 1]).



## Tykhonov for denoising

Tykhonov regularization with  $\lambda = 5$ .





TV-L2

$$E(u) = \frac{1}{2} \|\widetilde{u} - Au\|_2^2 + \lambda TV(u),$$

with  $TV(u) = \sum_{(i,j)} |(\nabla u)_{i,j}|$  the discrete version of the total variation.

Minimization : Convex, but not differentiable.

$$\frac{\partial E}{\partial u} = 2A^* (Au - \tilde{u}) - \lambda \operatorname{div} \frac{\nabla u}{\|\nabla u\|}.$$

Approximation of TV(u) by  $\int \sqrt{\|\nabla u\|^2 + \varepsilon^2}$ , with a small  $\varepsilon$  and gradient descent.

### TV-L2 for denoising

[Ruding-Osher-Fatemi, 1992] : denoising case, A = Id.

$$E(u) = \frac{1}{2} \|\tilde{u} - u\|_2^2 + \lambda T V(u).$$

### TV-L2 for denoising

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$$E(u) = \frac{1}{2} \|\tilde{u} - u\|_2^2 + \lambda TV(u)$$

Projection method using duality : [Chambolle 2004].

$$\begin{split} \hat{u} &= \arg\min_{u} E(u) \quad \Leftrightarrow \quad 0 \in \hat{u} - \tilde{u} + \lambda \partial TV(\hat{u}) \\ &\Leftrightarrow \quad \dots \text{ duality arguments} \\ &\Leftrightarrow \quad \frac{\tilde{u} - \hat{u}}{\lambda} = \Pi_{\kappa}(\frac{\tilde{u}}{\lambda}), \end{split}$$

with

$$\kappa = {\operatorname{div} p \mid \max_{x \in \Omega} |p(x)| \le 1}.$$

## TV L2 for denoising

A = Id,  $N \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma = 0.2$  (image range = [0, 1]).



# TV L2 for denoising

ROF, small  $\lambda$ .




# TV L2 for denoising

ROF, large  $\lambda$ .





#### TV-L1

$$E(u) = \frac{1}{2} \|\tilde{u} - Au\|_1 + \lambda TV(u).$$

- Contrast invariant
- Much more robust to outliers than TV-L2 (impulse noise, salt&pepper noise).

## Minimization?

TV-L1

Let  $X = \mathbb{R}^{n \times m}$ , it happens that

$$TV(u) = \max_{p \in X imes X} < 
abla u, p >_X - \iota_\kappa(p),$$

where

$$\iota_\kappa(p) = egin{cases} 0 & ext{if } p \in \kappa \ +\infty & ext{if } p 
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Thus,

$$\arg\min_{u\in X} \qquad \frac{1}{2}\|\tilde{u} - Au\|_{1} + \lambda TV(u) = \arg\min_{u\in X} \max_{p\in X\times X} \underbrace{\frac{1}{2}\|\tilde{u} - Au\|_{1}}_{G(u)} + \underbrace{\lambda < \nabla u, p >_{X}}_{} - \underbrace{\iota_{\kappa}(p)}_{F^{*}(p)}$$

[Chambolle-Pock, 2012] Primal-dual algorithm for (smooth or non smooth) convex problems of the form :

$$\min_{x \in X} \max_{y \in Y} \langle Kx, y \rangle + G(x) - F^*(y) \tag{1}$$

with

- X and Y finite-dimensional vector spaces
- $K : X \rightarrow Y$  linear
- $G: X \to \mathbb{R}^+$  and  $F^*: X \to \mathbb{R}^+$  convex, proper, lower-semicontinuous.

#### Chambolle-Pock primal-dual algorithm

#### Algorithm

- **1** Initialization : choose  $\tau$ ,  $\sigma > 0$ ,  $(x^0, y^0) \in X \times Y$
- **2** Iterations for  $n \ge 0$ :

$$y^{n+1} = \underbrace{\operatorname{prox}_{\sigma F^*}}_{backward step} \underbrace{(y^n + \sigma Kx^n)}_{forward step}$$
$$x^{n+1} = \operatorname{prox}_{\tau G}(x^n - \tau K^* y^{n+1}).$$

where  $prox_f$  is the proximity operator of a convex, l.s.c function f, defined as

$$\operatorname{prox}_{f}(x) = \arg\min_{z} f(z) + \frac{1}{2} ||x - z||^{2}.$$

Makes sense if  $prox_{\sigma F}$  and  $prox_{\tau G}$  are easy to compute (closed-forms).

## Chambolle-Pock algorithm for denoising with TV-L1

• 
$$G(u) = \frac{1}{2} \|\tilde{u} - u\|_1$$
  
•  $F^*(p) = \iota_{\kappa}(p).$   
•  $\langle Ku, p \rangle = \lambda \langle \nabla u, p \rangle$  (i.e.  $K = \lambda \nabla$  and  $K^* = -\lambda \operatorname{div}$ )

The proximal operators of G and  $F^*$  are

$$(\mathrm{prox}_{\tau G}(u))_{i,j} = \begin{cases} u_{i,j} - \tau & \text{if } u_{i,j} - g_{i,j} > \tau \\ u_{i,j} + \tau & \text{if } u_{i,j} - g_{i,j} < \tau \\ g_{i,j} & \text{if } |u_{i,j} - g_{i,j}| \le \tau \end{cases}$$
  
 $(\mathrm{prox}_{\sigma F^*}(p))_{i,j} = (\pi_{\kappa}(p))_{i,j} = rac{p_{i,j}}{\max(1, |p_{i,j}|)}.$ 

#### Chambolle-Pock algorithm for denoising with TV-L1

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$$(\operatorname{prox}_{\sigma F^*}(p))_{i,j} = (\pi_{\kappa}(p))_{i,j} = \frac{p_{i,j}}{\max(1, |p_{i,j}|)}.$$

- If  $G(u) = \frac{1}{2} \|\tilde{u} u\|_2^2$ , then  $(\operatorname{prox}_{\tau G}(u))_{i,j} = \frac{u_{i,j} + \tau \tilde{u}_{i,j}}{1 + \tau}$ , so the same algorithm can be applied to minimize TV-L2.
- If  $G(u) = \frac{1}{2} \|\tilde{u} Au\|_2^2$  with A a convolution matrix,  $\operatorname{prox}_{\tau G}(u)$  can also be computed explicitely.

## TV-L1 and TV-L2 for impulse noise

Impulse noise with p = 0.3



TV-L1 and TV-L2 for impulse noise

#### TV-L2





 $\mathsf{TV}\text{-}\mathsf{L1}$  and  $\mathsf{TV}\text{-}\mathsf{L2}$  for impulse noise

#### TV-L1





## TV-L2 for deconvolution

#### [Chambolle-Pock,2012] with A motion blur.



(a) Original image



(c) Wiener filter



(b) Degraded image



(d) TV-deconvolution

 $\min_{u_{\mid mask} = \tilde{u}_{\mid mask}} TV(u)$ 

$$\min_{\substack{u_{\mid mask} = \tilde{u}_{\mid mask}}} TV(u)$$
$$= \min_{u} TV(u) + \iota_{\mathcal{C}}(u) \text{ with } \mathcal{C} = \{u; u_{\mid mask} = \tilde{u}_{\mid mask}\}$$

``

- 1/



50% random missing pixels



TV inpainting



70% random missing pixels



TV inpainting



Original image

Troisième partie III

Multi-image restoration

Using several shots to increase image quality has become a common challenge in digital photography, movie post-production and remote sensing imaging.

#### A word on multi-image restoration

- denoising (burst denoising) in low light (avoid motion blur);
- Ø dynamic range increasing (HDR);
- 8 panoramas creation ;
- ø superresolution, 4k standard
- 6 color harmonization, style transfert





Take a burst of images  $U_1, U_2, \ldots, U_n$  with short-exposure times and average them after registration.

Assume an i.i.d. and centered additive noise of variance  $\sigma^2$ 

$$\forall i \in \{1,\ldots n\}, \quad \widetilde{U}_i = U_i + B_i.$$

For each pixel  $x \in \Omega$ ,

$$\frac{\widetilde{U}_1(x) + \ldots + \widetilde{U}_n(x)}{n} \xrightarrow[n \to +\infty]{}$$

and

$$\operatorname{Var}\left[rac{\widetilde{U}_1(x)+\cdots+\widetilde{U}_n(x)}{n}
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$$\operatorname{Var}\left[\frac{\widetilde{U}_1(x)+\cdots+\widetilde{U}_n(x)}{n}\right]=\frac{\sigma^2}{n}.$$

Fusing *n* images reduce the noise by a factor  $\sqrt{n}$ .

# A noisy image (Gaussian additive noise, $\sigma = 20$ )



# Mean of 5 images



# Mean of 10 images



# Mean of 20 images



# Mean of 40 images



# Original image



#### A real-life example of burst denoising



Figure 1.1: From left to right: (a) one long-exposure image (time=0.4 s, ISO=100), one of 16 short-exposure images (time=1/40 s, ISO=1600) and their average after registration. The long exposure image is blurry due to camera motion. (b) The middle short-exposure image is noisy. (c) The third image is about **four times** less noisy, being the result of averaging 16 short-exposure images. From [19].

From [Buadès et al., A note on multi-image denoising, 2009].

### Multi-image for denoising : impulse noise

#### Noise model

$$\forall x \in \Omega, \ \ \widetilde{U}_i(x) = (1 - \mathcal{T}_i(x)).U(x) + \mathcal{T}_i(x).W_i(x),$$

#### where

- $T_i \sim \text{Bernouilli of parameter } p$ ;
- *W<sub>i</sub>* uniform noise on [0, 255];
- $T_i$ ,  $W_i$ ,  $i \in \{1 \dots n\}$  are independent.

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How can we estimate U(x)?
## Impulse noise, one sample, p = 0.3



## Unbiased estimation from mean of 10 images



## Unbiased estimation from mean of 30 images



## Unbiased estimation from mean of 100 images



# Estimation from median, 10 images



# Estimation from median, 30 images



## Estimation from median, 100 images



What happens if we increase the noise level p? Is the median still a good estimator?

## Impulse noise, one sample, p = 0.6



# Estimation from median, 10 images



# Estimation from median, 30 images



## Estimation from median, 100 images



### Maximum-likelihood for impulse noise

The variables  $U_1(x), \ldots, U_n(x)$  follow the law

$$(1-p)\delta_{U(x)}+p\mathbf{1}_{[0,255]}.$$

Maximum likelihood

$$\begin{split} \widehat{U(x)} &= \arg \max_{U(x)} \mathbb{P}[\widetilde{U}_1(x), \dots, \widetilde{U}_n(x) | U(x)] \\ &= \arg \max_{U(x)} \sum_{k=1}^n \log \mathbb{P}[\widetilde{U}_k(x) | U(x)] \\ &= \arg \max_{U(x)} \sum_{k=1}^n \log \left[ (1-p) \delta_{U(x)}(U_k(x)) + p * \frac{1}{256} \right] \\ &= \arg \max_{U(x)} h(U(x)), \end{split}$$

with *h* the histogram of the values  $\{\widetilde{U}_1(x), \ldots, \widetilde{U}_n(x)\}$ .

## Estimation from histogram, 10 images, p = 0.6



## Estimation from histogram, 30 images, p = 0.6



## Estimation from histogram, 100 images, p = 0.6



#### [Haro, Buadès, Morel, 2012]



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#### General case

- Not always possible to take several images : astronomy, medical imaging, etc;
- Even if we can take several shots, a global registration is not always enough : object motions, large camera motions, etc.

SOLUTION?

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- Not always possible to take several images : astronomy, medical imaging, etc;
- Even if we can take several shots, a global registration is not always enough : object motions, large camera motions, etc.

#### SOLUTION?

Exploit the patch redundancy of natural images for restoration.

Quatrième partie IV

Patch-based methods

**Non local models** = all models (either variational, stochastic or geometric) which represent images by a set of local neighborhoods or *patches*, and make them collaborate regardless of their spatial position in the image.



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## Applications in image restoration and editing

#### Image editing and synthesis

- texture synthesis Efros-Leung [99]
- image *retargeting* or *reshuffling* Barnes et al. [09]
- style transfer Frigo et al. [16]

#### Image restoration

- denoising Buades et al. [05], Awate Whitaker [06], Dabov et al. [08], Lebrun et al. [12],
- non gaussien denoising, Poisson, Speckle Deledalle et al. [10], [12], impulse noise Delon Desolneux [13]
- inpainting Wexler et al. [04], Criminisi Perez [04], Newson et al. [14]
- interpolation Yu et al. [12], demosaicing Buades et al. [07]
- HDR Aguerrebere et al. [17], compression
- General inverse problems Peyré [08], PLE Yu et al. [12]



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## Texture synthesis : Efros-Leung (1999)

### Efros - Leung [99]

- Markov random fields models (inspired from Shannon models for text synthesis). We want to estimate  $p(u_i|x_i^u)$
- first paper with a patch-based approach : idea to exploit the patch *redundancy* in natural images
- global optimization instead of sequential synthesis Kwatra et al. [03]



### Texture synthesis : Efros-Leung (1999)





the following for all cooling reservances and the servance of it ndateears coune Tring rooms," as Heft he fast nd it l ers dat noears outseas ribed it last nt hest bedian Al. I econicalHomd ith Al. Heft ars of as da Lewindailf I lian Al Ths," as Lewing questies last aticarsticall. He is dian Al last fal counda Lew: at "this dailyears d ily edianicall. Hoorewing rooms," as House De fale f De und itical councestscribed it last fall. He fall. Hefft rs oroheoned it nd it he left a ringing questica Lewin. icars coecoms," astore years of Monica Lewinow see a Thas Fring roome stooniscat nowea re left a roouse bouestof MHe lelft a Lest fast ngine launesticars Hef ud it rip?" Triouself, a ring ind itsonestid it a ring que: astical cois ore years of Moung fall. He ribof Mouse re years ofanda Tripp?" That hedian Al Lest fasee yea ada Tripp?' Holitical comedian Alét he few se ring que olitical cone re years of the storears of as 1 Frat nica L ras Lew se lest a rime 1 He fas quest nging of, at beou

ut it becomes harder to lau ound itself, at "this dailyving zooms," as House Der rscribed it last fall. He fai it he left a ringing questio ore years of Monica Lewir inda Tripp?" That now seer ?olitical comedian Al Frar txt phase of the story will

Try by yourself! demo.ipol.im/demo/59/




Image denoising : NLMeans (2005)

Observation

$$\tilde{u} = u + n$$

with  $n \sim \mathcal{N}(0, \sigma^2)$ , we want to reconstruct u.

• Non Local Means, Buades, Coll, Morel, [05]

$$\forall i \in \Omega, \quad \mathsf{NLu}_i = \frac{\sum_{j \in \Omega} w_{i,j} \tilde{u}_j}{\sum_j w_{i,j}}.$$

with  $w_{i,j}$  weights measuring the similarity between patches centered at i and j, typically

$$w_{i,j} = e^{-\|x_i^{\tilde{u}} - x_j^{\tilde{u}}\|_2^2/2h^2}$$



Région uniforme

Région texturée

Contour géométrique

# Image denoising : NLMeans (2005) Buades, Coll, Morel, [05]



Noisy Image,  $\sigma=20$ 



NL-means

Try by yourself! demo.ipol.im/demo/bcm\_non\_local\_means\_denoising/

# Patch style transfer (2016)





Frigo et al. [16] MRF models

## Inverse problems

#### Model : observation v



Goal : estimate u from v



image u

noise

blur

missing data

## Inverse problems

Model : for each patch  $y_i$  from v

 $y_i = A_i \quad x_i + n_i$  observation acquisition operator unknown noise

Goal : estimate all clean patches  $x_i \in \mathbb{R}^p$  from the observations  $\{y_i\}_i$ 



image u

noise

blur

missing data

Assuming a prior distribution p(x) for X, the posterior distribution is

$$p(x|y) = \propto p(y|x)p(x) \propto e^{-\frac{\|Ax-y\|^2}{2\sigma^2}}p(x).$$

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#### Some restoration strategies

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For Gaussian priors, MAP = MMSE = Linear MMSE.

If  $X \sim \mathcal{N}(\mu, \Sigma)$  and  $N \sim \mathcal{N}(0, \sigma^2 I_p)$  are independent,

$$\hat{x} = \psi(y) := \underset{x}{\arg \max} \log p[x|y]$$

$$= \underset{x}{\arg \min} \frac{1}{2\sigma^2} (Ax - y)^t (Ax - y) + (x - \mu)^t \Sigma^{-1} (x - \mu)$$

$$= \mu + \Sigma A^t (A\Sigma A^t + \sigma^2 I_p)^{-1} (y - A\mu)$$

## DENOISING WITH GAUSSIAN PRIORS

Bayesian framework with Gaussian or GMM priors on patches, EPLL [11], NL-Bayes [12], PLE [12], S-PLE [13], DA3D [15].



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# WHY GAUSSIAN OR GMM PRIORS?



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## DENOISING WITH THE "RIGHT" MODEL



How to restore the image from the set of restored patches?
How to estimate (μ, Σ) from the degraded patches {y<sub>i</sub>}?

# Reconstruction of u from restored patches



Central value

# Reconstruction of u from restored patches



#### Central value

Aggregation of estimators

# Reconstruction of u from restored patches



Central value

Aggregation of estimators

Global Optimisation in u, EPLL, Zoran-Weiss [11]

$$rgmin_u rac{\lambda}{2} \|Au - v\|_2^2 - \sum_j \log p(x_i^u)$$

# How to infer Gaussian or GMM priors?

- Global or spatially local models, Deledalle et al. [11]
- Local Gaussian models : nearest neighbours, NL-Bayes denoising, Lebrun et al. [13]
- GMM  $\sum_{k=1}^{K} \pi_k \mathcal{N}(\mu_k, \Sigma_k)$ 
  - estimated on a large external database, EPLL, Zoran-Weiss [11]
  - estimated on the degraded image with a synthetic initialization, PLE, Yu et al. [12];
  - estimated on the degraded image, Teodoro et al. [15];
  - estimated on the degraded image, based on mixture of PPCA, SURE-PLE, Wang et Morel [13]
  - estimated on the degraded image, based on HDDC, HDMI (High-Dimensional Mixture Model for Image denoising), Houdard et al. [17]
- GGMM (generalized Gaussian mixture models) Deledalle et al. [18]

## The curse of dimensionality



Estimation of **Sample Covariance Matrices**  $\widehat{\Sigma}$  from *n* samples in high dimension is difficult : estimates tend to be ill-conditionned or even singular...

 $\rightarrow$  but  $\widehat{\Sigma}$  has to be inverted to compute Wiener estimators...

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 $\rightarrow$  but  $\widehat{\Sigma}$  has to be inverted to compute Wiener estimators...

#### Some workarounds

- use small patches + flat area trick (3 × 3 or 5 × 5 in NL-Bayes, Lebrun et al. [13])
- use covariances of fixed lower dimensions, SURE-PLE, Wang et al. [13]
- add regularization  $(+|\varepsilon|I_p)$  or hyperpriors HBE, Aguerrebere et al. [17]
- infer a specific dimension for each Gaussian HDMI, Houdard et al. [17]

# HDMI (BOUVEYRON, D., HOUDARD [17])

Assume that patches live in low-dimensional subspaces, specific to their latent groups.



 $Z \in \{1 \dots K\}$  latent r.v. indicating the group from which X is generated

 $X_{|Z=k} \sim \mathcal{N}\left(\mu_k, U_k \Lambda_k U_k^t\right),$ 

with

- $\mathcal{P}[Z=k]=\pi_k$
- $U_k p \times d_k$  orthonormal,
- $\Lambda_k = \operatorname{diag}(\lambda_{k1}, \ldots, \lambda_{kd_k}),$

•  $\mu_k \in \mathbb{R}^p$ 

This model

- is a generalization of the full GMM if  $d_k = p$ ,
- has strong links with the MPPCA model, Tipping, [96].

## The HDMI model

The distribution of Y is also a mixture of Gaussians :

$$p(y) = \sum_{k=1}^{K} \pi_k \mathcal{N}(y; \mu_k, \Sigma_k)$$
 with  
 $\Sigma_k = U_k \Lambda_k U_k^t + \sigma^2 I_p.$ 

Let  $Q_k = [U_k, R_k]$  be a  $p \times p$  matrix made of  $U_k$  and an orthonormal complementary, then

$$\Delta_{k} = Q_{k}^{t} \Sigma_{k} Q_{k} = \begin{pmatrix} \begin{array}{ccc} a_{k1} & 0 & & \\ & \ddots & & \\ 0 & & a_{kd_{k}} & & \\ & & & \sigma^{2} & 0 \\ & 0 & & \ddots & \\ & & & 0 & \sigma^{2} \end{pmatrix} \end{pmatrix} \begin{pmatrix} a_{k} & & \\ b_{k} & & \\ b_{k} & & \\ c_{k} & c_{k} & \\ c_{k} & \\$$

with  $a_{kj} = \lambda_{kj} + \sigma^2$  and  $a_{kj} > \sigma^2$ , for  $j = 1, ..., d_k$  and  $k = 1, \ldots, K$ .

## DENOISING WITH THE HDMI MODEL

To denoise the patch  $y_i$ , we compute

$$\hat{x}_i = \mathbb{E}[X|Y = y_i].$$

**Proposition.** Under the assumptions of the HDMI model, the conditional expectation  $\mathbb{E}[X|Y = y_i]$  can be written

$$\mathbb{E}[\mathbf{X}|\mathbf{Y}=\mathbf{y}_i] = \sum_{k=1}^{K} \mathcal{P}(Z=k|\mathbf{Y}=\mathbf{y}_i)\psi_k(\mathbf{y}_i),$$

where

$$\psi_k(\mathbf{y}) = \mu_k + (\Sigma_k - \sigma^2 \mathbf{I}_p) \Sigma_k^{-1} (\mathbf{y} - \mu_k)$$
  
=  $\mu_k + \tilde{Q}_k \left( \mathbf{I}_p - \sigma^2 \Delta_k^{-1} \right) \tilde{Q}_k^t (\mathbf{y} - \mu_k),$ 

with  $\tilde{Q}_k = [U_k, 0_{p,p-d_k}].$ 

Before denoising the patches  $\{y_1, ..., y_n\}$ , the HDMI model has to be inferred from the data :

- estimate model parameters  $\theta = \{\pi_k, \mu_k, a_{kj}, \sigma^2, Q_k\},\$
- determine hyper-parameters K and  $d_k$ .

## Model inference

**EM** algorithm : maximize *w.r.t.*  $\theta$  the conditional expectation of the complete log-likelihood :

$$\Psi(\theta, \theta^*) \stackrel{\text{def}}{=} \sum_{k=1}^{K} \sum_{i=1}^{n} t_{ik} \log \left( \pi_k p\left( y_i; \theta_k \right) \right),$$

where  $t_{ik} = \mathbb{E}\left[Z = k | y_i, \theta^*\right]$  and  $\theta^*$  a given set of parameters.

- E-step estimation of t<sub>ik</sub> knowing the current parameters
- M-step compute maximum likelihood estimators (MLE)

$$\widehat{\pi}_k = \frac{n_k}{n}, \qquad \widehat{\mu}_k = \frac{1}{n_k} \sum_i t_{ik} y_i, \qquad \widehat{S}_k = \frac{1}{n_k} \sum_i t_{ik} (y_i - \mu_k) (y_i - \mu_k)^T,$$

with  $n_k = \sum_i t_{ik}$ . Then  $\widehat{Q}_k$  is formed by the  $d_k$  first eigenvectors of  $\widehat{S}_k$  and  $\widehat{a}_{kj}$  is the *j*th eigenvalue of  $\widehat{S}_k$ .

The hyper-parameters K and  $d_1, \ldots, d_K$  cannot be determined by maximizing the log-likelihood since they control the model complexity.

We propose to set K at a given value (for instance K = 90) and to choose the intrinsic dimensions  $d_k$ :

- using an heuristic that links  $d_k$  with the noise variance  $\sigma$  when it is known (supervised case);
- using a model selection tool in order to select the best  $\sigma$  when unknown (unsupervised case).

# ESTIMATION OF INTRINSIC DIMENSIONS when $\sigma$ is known

**Heuristic.** Given a value of  $\sigma^2$  and for k = 1, ..., K, we estimate the dimension  $d_k$  by

$$\widehat{d}_k = \operatorname{argmin}_d \left| \frac{1}{p-d} \sum_{j=d+1}^p \widehat{a}_{kj} - \sigma^2 \right|$$

Each value of  $\sigma$  yields a different model, we propose to select the one with the better BIC (Bayesian Information Criterion)

$$\operatorname{BIC}(\mathcal{M}) = \ell(\hat{\theta}) - \frac{\xi(\mathcal{M})}{2}\log(n),$$

where  $\xi(\mathcal{M})$  is the complexity of the model.

# Role of the intrinsic dimensions $d_k$



**FIGURE** – Map of intrinsic dimensions  $d_k$ .

## EFFECT OF THE DIMENSION REDUCTION INSIDE EM



# NUMERICAL EXPERIMENTS

Original Image



# NUMERICAL EXPERIMENTS

Noisy image  $\sigma = 50$ 



## NUMERICAL EXPERIMENTS

Denoised with BM3D, Foi et al. 2007,  $\mathsf{psnr}=27.17\mathsf{dB}$ 


Denoised with FFDNet, Zhang et al. 2018, psnr = 27.58dB



Denoised with HDMI<sub>sup</sub> K = 90, psnr = 27.28dB



Original Image



Noisy image  $\sigma = 50$ 



Denoised with BM3D, Foi et al. 2007, psnr = 26.55.dB



Denoised with FFDNet, Zhang et al. 2018, psnr = 27.45dB



Denoised with HDMI<sub>sup</sub> K = 90, psnr = 27.05dB





Best of both worlds, psnr = 27.86dB



Original Image

